

Partisan Dislocation: A Precinct-Level Measure of Representation and Gerrymandering

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Abstract

We introduce a fine-grained measure of the extent to which electoral districts combine and split local communities of co-partisans in unnatural ways. Our indicator – which we term *Partisan Dislocation* – is a measure of the difference between the partisan composition of a voter’s geographic nearest neighbors and that of her assigned district. We show that our measure is a good local and global indicator of district manipulation, easily identifying instances in which districts carve up clusters of co-partisans (cracking) or combine them in unnatural ways (packing). We demonstrate that our measure is related to but distinct from other approaches to the measurement of gerrymandering, and has some clear advantages, above all as a complement to simulation-based approaches. It can also be used prospectively by district-drawers who wish to create maps that reflect voter geography, but according to our analysis, that goal is sometimes in conflict with the goal of partisan fairness.

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1 Introduction

In an era of partisan polarization, opposition to gerrymandering is a rare instance of bipartisan consensus among voters. While some opponents of gerrymandering are primarily motivated by perceived unfairness in the transformation of votes to seats for their preferred party, revulsion of the practice runs deeper. Even in states where the Republican candidates are the beneficiaries, for example, clear majorities of Republican voters have advocated anti-gerrymandering provisions both in surveys and referendums. Many voters are motivated by the notion that they—along with geographic clusters of like-minded neighbors—should elect representatives who can advocate for them in the state capital or in Washington. What rankles is when, in order to increase its seat share or harm an enemy, the incumbent party breaks up such neighborhoods and combines fragments of disparate ones that have little in common.

For those who see value in a system of political representation based on small geographic districts, much of the value lies in allowing neighbors who live in the same community, and hence share common interests and concerns, to be represented by a single politician. In other words, a perceived danger of gerrymandering is not just that it leads to *global* unfairness in the transformation of votes to seats in a U.S. state, but that it leads to an abridgment of *local* rights of representation. Justice Roberts articulated this view when writing for the majority in *Gill v. Whitford* about the issue of legal standing to sue: “[t]o the extent the plaintiffs’ alleged harm is the dilution of their votes, that injury is district specific. [...] In this gerrymandering context that burden arises through a voter’s placement in a “cracked” or “packed” district.”

In this paper, we introduce a new measure of cracking and packing that is completely divorced from concerns about what is the “fair” share of seats that a party should receive when it obtains a specific share of the vote. We demonstrate that it is possible to clearly identify a partisan gerrymander without making normative claims about how many seats a particular party “deserves,” and without referencing seat shares at all. Rather, we measure what we call *Partisan Dislocation*—the extent to which a redistricting plan unnaturally separates individuals from local communities of co-partisans.

Our goal in developing this measure is to add something distinctive to the growing statistical toolkit used to identify partisan manipulation in the redistricting process. We are able to avoid some of the assumptions, controversies, and computational demands associated with existing approaches, most of which conflate the concepts of global partisan fairness and gerrymandering. Our measure also allows us to identify which *individual neighborhoods* have been packed or cracked in the creation of *individual districts*, but these measures can also be aggregated to the level of districts or states to measure the overall level of gerrymandering.

It might seem at first blush that identifying packed or cracked districts can be accomplished simply by looking at their partisan composition: one party will have a very high vote share in a packed district, and one party’s vote share will be just below 50%) in a cracked district. But partisan composition turns out to be an insufficient statistic for this task because partisan geographic clustering — for example, that of

Democrats in cities — may naturally give rise to districts in which one party has a very high vote share, not because of the political machinations of district architects, but instead because the party’s members live in close proximity to one another. Similarly, if a party receives 45% of the vote in a district drawn by its opponent, its supporters may have been intentionally cracked, but it could just as well be the case that there were too few of them in that region of the state to form a majority.

To address this challenge, we present a measure of the degree to which a representative individual voter is the victim of packing or cracking. In particular, we examine the degree to which the *partisan composition* of a voter’s actual electoral district differs from the partisan composition of their geographic neighborhood. Where these measures differ dramatically — where, for example, a voter whose k nearest neighbors (where k is the number of people in the voter’s actual legislative district) are mostly Democrats, but despite this their district is mostly Republican — we term that voter *partisan dislocated*.

As we will show in Section 4, *Partisan Dislocation* turns out to be a very good systematic measure of packing and cracking. Areas where voters are dislocated— that is, where they find themselves in districts with substantially different political compositions than their geographic neighborhoods—are very often in districts in which voters have been carefully carved out of their more natural communities (i.e. they have been “cracked” or “packed”) for electoral advantage. Moreover, our measure does *not* identify “naturally packed” districts as gerrymanders, such as those emerging in the core of large, highly Democratic cities, where districts inevitably have large vote shares for a single party due to residential partisan clustering. In such cases, the partisan composition of the district is often consistent with that of the voter’s geographic neighborhoods. As we discuss in later sections, this results in a measure that tends to track with current jurisprudence about what constitutes a gerrymander, though it may not be satisfying to those who dispute the emerging normative rationale for legal gerrymandering standards.

Next, we attempt to validate the aggregate statewide dislocation score as a global measure of gerrymandering by comparing it with some of the other measures that have become dominant both in the academic literature and in the courts. First, we discover that when focusing on enacted districting plans, there is a reasonably high correlation between global Partisan Dislocation and simple global measures of partisan fairness, like the mean-median difference in vote shares. Second, following the practice that has become common in court cases, we create a large ensemble of simulated redistricting plans for each state, and calculate the difference between the mean Democratic seat share in the ensemble and the Democratic seat share associated with the enacted plan. We find that this gap is highly correlated with the average absolute value of Partisan Dislocation across all voters. Likewise, we find high correlations between dislocation and other proposed measures of partisan gerrymandering that focus on the relationship between votes and seats.

However, our measure also captures something distinctive. Some clear efforts at packing and cracking are not picked up by existing global approaches to votes and seats,

in part because these are insufficiently sensitive to factors like incumbency, variation in the spatial distribution of support from one election to another, and efforts to pair incumbents or harm specific enemies. Partisan dislocation, by contrast, is well suited to identifying these forms of manipulation that generate harms that are hard to detect through global measures.

We show that Partisan Dislocation is useful not only as a simpler, far less computationally intensive alternative to computer simulations that requires fewer assumptions, but more importantly, as a complement to the simulation approach. Once one has generated 100,000 redistricting plans, it is not always clear what to do with them. We show that some of the most gerrymandered states are those where the global dislocation score of the enacted plan is a clear outlier relative to the the distribution of these scores in the redistricting ensemble. In some cases, this is much harder to see using the traditional comparison of anticipated seats in the enacted and simulated plans. Thus, our measure is a valuable metric on which to compare enacted and simulated plans.

While simulation-based methods can be used to identify gerrymanders *ex post*, they offer less guidance to mapmakers who might wish to draw districts that keep local clusters of co-partisans together. By providing a localized indicator of which specific precincts are dislocated in a specific plan under consideration, our measure could be useful in the redistricting process. For instance, compliance with the Voting Rights Act will often require significant partisan dislocation. However, when trying to draw districts in a specific region so as to make sure minorities can elect candidates of choice, among a variety of alternatives, some will generate much higher levels of partisan dislocation than others. Dislocation measures can therefore be used to help planners pick the least disruptive methods of achieving other objectives.

An advantage of partisan dislocation is that it is a relatively “pure” measure of gerrymandering that is distinctive from prevailing notions of overall fairness like partisan symmetry (Katz, King and Rosenblatt 2020). This allows us to begin exploring the relationship between two rather distinctive normative goals that might motivate those drawing electoral districts. Reformers often assume that by minimizing gerrymandering, they will also facilitate partisan symmetry. We demonstrate that this is very often not the case. We pay special attention to ensembles of redistricting plans that minimize (and maximize) dislocation. In some states—especially those of the 19th century manufacturing core of the Northeast and Upper Midwest, where Democrats are highly concentrated in space—we observe that the redistricting plans that minimize dislocation are characterized by high levels of partisan asymmetry. In these states, maps that keep partisan neighborhoods together will produce transformations of votes to seats that advocates of partisan symmetry would consider unfair. The goal of keeping communities of like-minded neighbors together will often be in conflict with the goal of promoting partisan fairness.

2 Partisan Fairness versus Gerrymandering

Gerrymandering is often viewed as unfair because it allows a party to achieve a seat share far beyond its vote share, or in some conceptualizations, beyond the seat share that would have been obtained with a non-partisan redistricting process. In the most obvious normative failure, a party with less than half of the statewide votes can receive more than half of the seats, which happens routinely in U.S. state legislatures. This is a global notion of representational harm, driven by the intuitive notion that the statewide vote-seat curve in a two-party system should be symmetric in its treatment of both parties. In this view of representation, courts should be suspicious of asymmetries in the transformation of votes to seats, and redistricting bodies should explicitly seek to draw symmetric plans.

Federal courts have expressed skepticism of the notion that the U.S. Constitution requires partisan symmetry, and have been reluctant to accept a role in measuring or enforcing it. It is clear that asymmetries can emerge in the transformation of votes to seats due to the geographic arrangement of partisans, even if the districts were drawn without partisan intent (Chen and Rodden 2013; Gudgin and Taylor 1979). For instance, in an evenly divided state, a party with a highly concentrated support base might end up with substantially less than half of the seats because it runs up large surpluses in core support areas where its voters are “packed”—e.g. Democratic candidates in large cities—while losing by smaller margins in the pivotal districts where its supporters are “cracked” as a result of residential patterns and the historical development of the party system. It seems unlikely that federal courts would be willing to strike down a map where partisan asymmetry cannot be clearly linked with intentional decisions of line-drawers. Thus in the context of gerrymandering litigation, the terms packing and cracking imply *partisan intent*.

In order to establish this type of intent, plaintiffs have developed a variety of techniques to sample from the very large number of potential alternative redistricting plans, with the goal of demonstrating that the partisanship of the enacted map was an extreme outlier relative to the ensemble of sampled maps and is thus unlikely to have emerged without significant effort on the part of mapmakers. For overviews of these techniques, see Chen and Rodden (2015); Cho and Liu (2016); Magleby and Mosesson (2018); Mathematicians’ Amicus Brief (2018); Mattingly and Vaughn (2014); Pegden (2017); Pegden, Rodden and Wang (2018).

This approach has been successful, but it is not without challenges. First, there are a variety of alternative techniques for sampling from the vast number of alternative plans. Some approaches are likely to sample only relatively compact plans, while others sample a much broader range of possible plans, with implications for whether specific plans under evaluation might end up being designated as outliers. There is no obvious way to decide which ensembles of plans are the “correct” baseline Best et al. (2018). Moreover, debates over these issues are highly technical (as they often relate to the exact acceptance probability parameters in Markov Chain Monte Carlo simulations), and thus difficult to present to non-specialist audiences like politicians and judges.

Another challenge is deciding what to do with an ensemble of alternative plans once one has generated it. What technique should one use for characterizing the partisanship of each district? Which precinct-level election results should be considered? What if presidential and attorney general elections lead to different inferences? Should some kind of swing, perturbation, or other hypothetical alternative election outcome be considered? Should the partisanship of each district be determined according to a discrete cut-point, or should one consider probabilities of victory for each party in each hypothetical district, perhaps based on an empirical model? These decisions are quite consequential in practice, and even in highly gerrymandered states, there are usually specific election results, or plausible-sounding ways of applying the uniform swing, that will make the hypothetical seat shares associated with the simulated and enacted plans appear to be similar.

Furthermore, measures of gerrymandering that focus on anticipated seats might miss some of the subtleties of the art. Safe seats occupied by popular incumbents, for instance, might be misclassified as losses for the incumbent party when using statewide or presidential results to classify seats. Sometimes the goal of gerrymandering is to force incumbents of the out-party to run against one another, to oust specific representatives of the out-party, to help a member of the in-party recover from a scandal, or even to harm renegade members of the in-party.

It would therefore be helpful to have an alternative measure of intentional gerrymandering that sidesteps some of the controversies about sampling. Furthermore, it would be useful to have a metric, other than the hypothetical seat shares of the parties, or some transformation of that quantity, along which to compare an ensemble of sampled plans with the plan that is being evaluated as a potential gerrymander.

Moreover, concerns about *global* representational harm to a political party are not the only basis for concern about gerrymandering. Fundamental to a political system featuring single member districts is the idea that there is value in voters from the same community who live in the same area being represented by a single politician. Arguments for this are multifaceted — voters in the same neighborhood are likely to share political interests; voters in the same area are better able to communicate and coordinate with one another; politicians can better maintain connections with voters in the same area; voters in the same area are especially likely to belong to the same social communities — but all suggest the importance of voters being located in districts with their geographic peers. For many voters, the reality falls far short of this ideal. Instead, efforts to gerrymander districts for political purposes result in clusters of voters being carved out of their natural communities and pooled with other voters in an effort to dilute their political influence. This may not only undermine the political effectiveness of these voters, but it may also deprive them of the benefits associated with belonging to a coherent constituency.

Yet existing global measures of gerrymandering focus exclusively on votes and seats, and are thus poorly suited to identifying deviations from this ideal. This is a significant weakness. A recent empirical study indicates that “packed” and “cracked” voters might receive fewer fiscal transfers (Stashko N.d.). Another empirical study suggests that

cracking and packing of like-minded communities is associated with voters who are less engaged in politics, and politicians who provide inferior representation (Stephanopoulos 2012).

This is the notion of representational harm articulated by Justice Roberts in *Gill*, and it seems likely that this is the notion that motivates the opprobrium of gerrymandering among many Americans—even those whose favored party might benefit from it. It is also plausible that some state courts could adopt Roberts’ notion of representational harm. Thus, it is worthwhile to develop a measure of gerrymandering that corresponds to this notion of harm.

Finally, it is possible that a future Supreme Court majority will rule that state legislatures do not have the authority to delegate the task of district-drawing to independent commissions, even in response to overwhelming majorities of voters in referendums (though of course whether independent commissions are a panacea remains a topic of debate (Henderson, Hamel and Goldzimer 2018)). In that event, the only viable way to curb partisan gerrymandering would be through reforms like that implemented in Florida, where the legislature is still tasked with the job of drawing legislative district boundaries, but is forbidden from considering partisanship when doing so. In order to hold legislators accountable, it may be helpful to establish an empirical indicator of intentional packing and cracking.

3 Measuring Partisan Dislocation

In this section, we formally introduce a measure designed to meet this goal: *Partisan Dislocation*. In simple terms, Partisan Dislocation is a measure the difference between the partisan composition of a voter’s geographic nearest neighbors and the partisan composition of the district to which they have been assigned. More formally, for a voter v in district d as:

$$dislocation_v = dem_vote_share_d - \sum_{n \in N_v} \frac{\mathbb{1}_{n \text{ is Democrat}}}{|N|} \quad (1)$$

Where N_v is the set of the k nearest neighbors of voter v , where k is the average number of people in the relevant electoral districts. Large positive values indicate individuals whose district is substantially more Democratic than their nearest neighbors, while large negative values are indicative of individuals in districts that are substantially more Republican than their nearest neighbors.

Data and Estimation

For this paper, Partisan Dislocation is computed using precinct boundary files and electoral returns from the 2008 Presidential Election. We chose this election because presidential elections ensure that in our cross-sectional analyses across states, all voters

are considering the same slate of candidates, and because 2008 is the most recent Presidential Election for which precinct-level boundary files and returns are available for all 49 states that use precincts.¹

This data is used to calculate Partisan Dislocation as follows:²

1. First, representative voter points are generated in each precinct in proportion to the number of Democratic and Republican votes recorded.³ For example, in a precinct with 100 votes for Obama and 50 for McCain, we would generate (in expectation) twice as many representative Democratic voter points as Republican voter points. That precinct would also have twice as many total representative voter points (in expectation) as a precinct with 50 votes for Obama and 25 votes for McCain. We use representative points (rather than creating one point for every vote cast) for computational tractability.
2. Each voter point is placed uniformly at random within the boundaries of each precinct. This generates a distribution of representative voter points across the entire United States that closely mirrors the true distribution of voters (we discuss deviations from the true voter distribution due to sampling error and placing voters uniformly-at-random placement within precincts below).
3. For each voter point v , we then calculate the share of that v 's k nearest neighbors who represent Democratic voters. This is our estimate for the partisan composition of v 's geographic neighbors.
 - Note that the value of k is selected so that the number of neighbors considered represents the average number of voters in a single electoral district. As a result, this number varies by the legislative districts being studied. For US House districts, for example, the value of k used ensures that the number of neighbors considered represents 700,000 real voters. For the California upper legislative chamber, by contrast, k is chosen to represent the number of voters in the average California upper legislative district ($\sim 300,000$).
4. For direct comparability, the partisan composition of each representative voter v 's actual 2014 electoral district is then calculated as the Obama share of votes cast for Obama or McCain at precincts within that district.
5. Finally, the *Partisan Dislocation* score for each representative voter v is calculated by subtracting the Democratic vote share of v 's k nearest neighbors from the Democratic vote share of v 's district.⁴

¹Oregon does not record precinct-level results due to its vote-by-mail system.

²A pip-installable package for implementation of this strategy is currently under development, and will soon be available at http://www.github.com/nickeubank/partisan_dislocation.

³We do not consider third parties in this analysis.

⁴Note that because any uniform swing – adding a constant value to one party's vote share in all precincts to adjust for the relative popularity of a candidate (e.g. subtracting 3.69% to 2008

Measurement Error

Our use of 2008 Presidential Election precinct returns results in two forms of measurement error: sampling error from the use of representative voter points (instead of one point per vote cast) and random placement within districts, and spatial error from distributing our representative voter points *uniformly* within the boundaries of each precinct (since real voters are not usually uniformly distributed within precincts).

The first of these – sampling error – is relatively easy to quantify. As detailed in Appendix A, repeatedly re-generating our representative voters (which incorporates both re-sampling the number of voter points per precinct *and* random placement within each precinct) causes very little variation in resulting Partisan Dislocation scores.⁵⁶

The second source of error – error due to our uniformity assumption – is harder to quantify. However, most precincts are *very* small in proportion to the electoral districts being analyzed, as a result of which the space for error within each precinct is quite small in proportion to the geographic scale of the districts (or the area over which the corresponding number of nearest neighbors reside). It is worth noting, however, that the relative size of precincts (and thus the relative size of potential placement errors) is greater for smaller electoral districts (e.g. lower state legislative electoral districts) than for larger districts (e.g. US House districts). As such, this source of error is of greater concern as one applies these methods to smaller and smaller scales. As a result, this approach may not be appropriate for, say, city council districting analyses.

4 Partisan Dislocation, Packing and Cracking

In this section, we demonstrate the ability of this measure to detect incidents of deliberate packing and cracking, a task that is perhaps best illustrated by mapping out the distribution of partisan-dislocated voters in several states.

We begin by examining two of the most clear-cut cases of packing and cracking in the United States – the US House of Representatives electoral districts built around Austin, Texas (a clear case of cracking) and the US House of Representatives districts formed out of Baton Rouge and New Orleans in Louisiana (a clear case of packing). These two cases are illustrated below in Figures 1 and 2. Voters colored red are those

Presidential returns to adjust for the fact that Obama won 53.69% of the two-party vote in the US) – would be applied equally to both calculating of the partisan composition of voters’ nearest neighbors and their district partisan composition, this measure is uniform-swing-invariant.

⁵It is computationally intractable to draw enough samples to precisely estimate of the variance introduced via this bootstrapping method.

⁶The number of representative voter points we generate in each precinct for each party is determined by taking a binomial draw from the total number of actual voters. The binomial probability varies by state-chamber, but is equal to $prob_k = \frac{\text{numberofdistricts}}{\text{numberofvotersinstate}} * k$, where $k=1,000$ for state legislative districts and 5,000 for US Congressional districts. This probability generates k voters per district in expectation. A larger number of representative points per district are used for US Congressional districts to adjust for the fact that the larger size of US Congressional districts results in a lower binomial sampling probability per precinct for a given target k , increasing the sampling variance.

who have been assigned to an electoral district that is substantially more Republican than their nearest neighbors, while voters colored blue are assigned to districts that are substantially more Democratic than their nearest neighbors. Lighter colors indicate voters for whom the difference between the partisanship of the voter’s district and her nearest neighbors is small, while darker colors indicate greater dislocation. Note that the colors are unrelated to the *partisanship of the individual voter* – they reflect only the difference between the voter’s community and that of her district.

In Figure 1, it is clear to see how Austin has been effectively cracked into a set of pizza-wedge shaped districts, each of which grabs a portion of the (largely Democratic) residents of Austin and pools them with a rural population of Republicans to create Republican-majority districts. This cracking is evident in the high dislocation scores for residents of Austin, who live in highly Democratic communities but have nevertheless been carved up and placed in Republican districts. The lone exception to this pattern is the long, narrow district that pools a small collection of Austin voters with Democrats in San Antonio to create a packed district, a form of manipulation which is evident in the high dislocation scores of the voters in the middle of this long, narrow district – voters in rural Republican communities who these contorted districts have dislocated in order to make this pooled district.

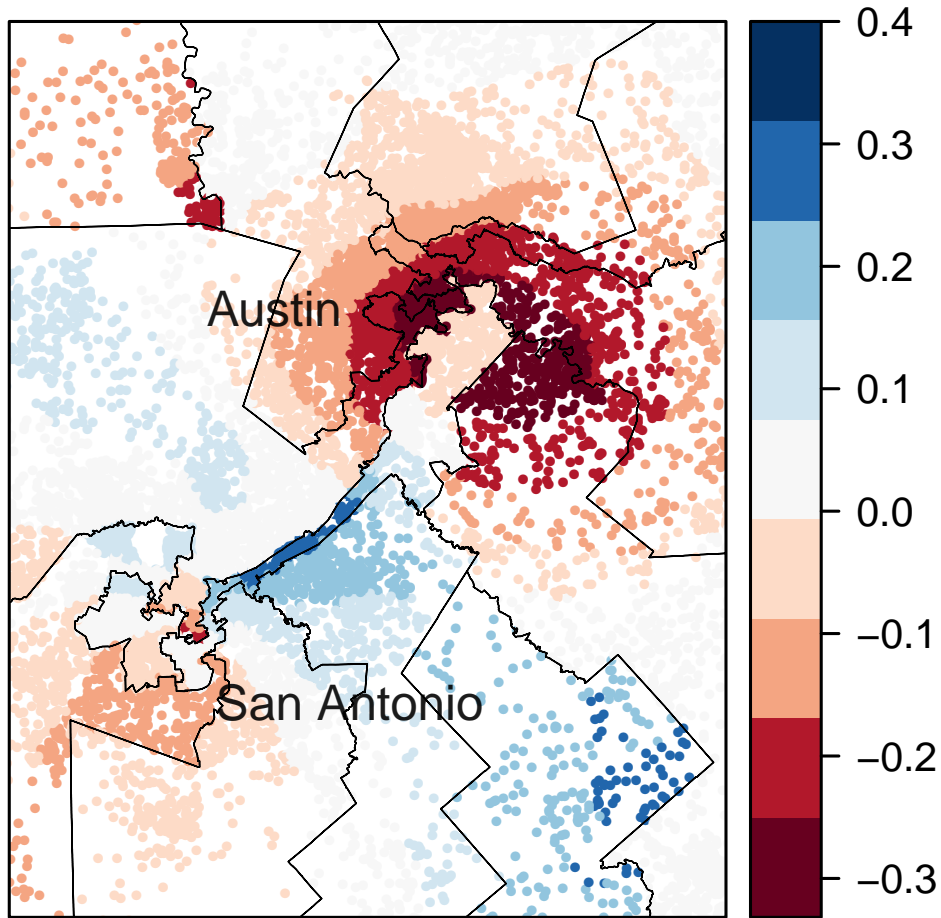
In Figure 2, we see an illustration of extreme packing in the district that pulls together New Orleans and Baton Rouge. Here we see that voters in both Baton Rouge and New Orleans have been placed in a district that is dramatically more Democratic than their local communities (as shown by regions of bright blue in both cities). At the same time, there is also evidence of cracking in the northern portion of New Orleans which has been carved away from the rest of the city and pooled with (more Republican) voters on the other side of Lake Pontchartrain.

The cases of Baton Rouge and New Orleans also make it clear that while Partisan Dislocation is a strong indicator of deliberate district manipulation, it cannot speak to whether that manipulation is normatively desirable. In the case of Baton Rouge and New Orleans, for example, part of the rationale for this district is an effort to create a majority-minority district in order to comply with the Voting Rights Act.

With that said, what the Partisan Dislocation measure *can* do is evaluate whether majority-minority districts like the Louisiana 2nd district have been drawn in a manner that *minimizes* overall dislocation. As such, Partisan Dislocation offers a method for comparing proposals for potential majority-minority districts in a way that makes it possible to police the potential abuse of the majority-minority imperatives for political advantage. After conducting analysis to ascertain the desired racial characteristics of majority-minority districts, it is possible to contrast dislocation scores among a variety of plans with the desired characteristics. This can be done not just for the entire map, but for specific areas in the vicinity of districts designed to facilitate minority representation. Indeed, as we show in Section 5, majority-minority districts can be achieved in Louisiana while also achieving lower levels of voter dislocation than in this enacted plan, which creates a majority-minority district with a black share of the voting-age population that is well beyond anything that could plausibly be required by

Figure 1: Partisan Dislocation in Austin, Texas US House Districts

Austin & San Antonio

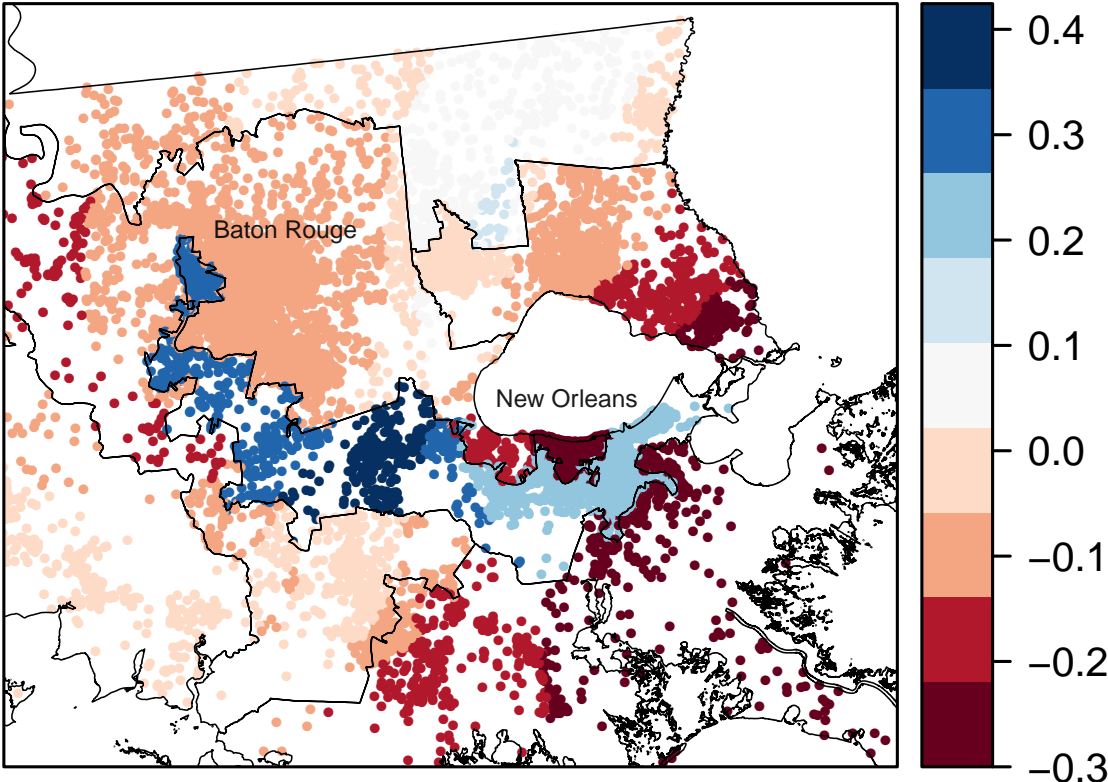


District D–Share minus Voter's KNN D–Share
State Avg Abs. Dislocation: 0.077

Notes: The above maps plot Partisan Dislocation scores for a set of representative voters. Dislocation is calculated as the difference in the Democratic vote share of each voter's assigned district and the Democratic vote share of her k nearest neighbors, where k is the average number of people assigned to each electoral district. District vote shares and the partisanship of nearest neighbors are estimated using precinct-level 2008 US Presidential vote shares as detailed in Section 3. Actual 2014 electoral district boundaries are also included.

Figure 2: Partisan Dislocation in New Orleans and Baton Rouge, Louisiana US House Districts

Baton Rouge & New Orleans, 2010–2019



District D–Share minus Voter's KNN D–Share
State Avg Abs. Dislocation: 0.102
Dist 2 Avg Absolute Dislocation: 0.236

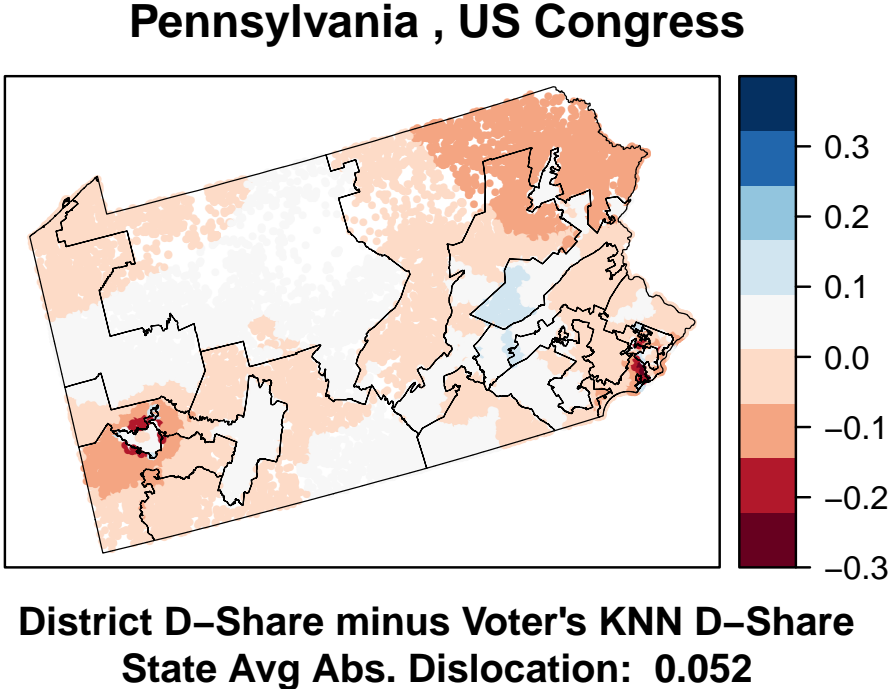
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the Voting Rights Act via extreme district manipulation.

Looking at these figures, one might worry that dislocation is simply a proxy for district compactness. However, this is not the case. Not only are the measures theoretically distinct — one could draw a district with arbitrarily low or high compactness in a state where voters are uniformly distributed, and dislocation would always remain zero — but as discussed in Appendix C, they are also quite empirically distinct; more compact districts do tend to have lower levels of dislocation, but the correlation is only ~ 0.275 .

While especially illustrative, these extreme examples are far from unique. Next, let us consider the state of Pennsylvania, a subject of extensive gerrymandering litigation. Figure 3 maps voter Partisan Dislocation for a representative set of voters. Note that similar patterns can be seen in a number of states who have been accused of gerrymandering in recent years. See Appendix B for analogous maps of North Carolina, Texas, Louisiana, and Maryland.

Figure 3: Partisan Dislocation in Pennsylvania US House Districts



Notes: The above maps plot Partisan Dislocation scores for a set of representative voters. Dislocation is calculated as the difference in the Democratic vote share of each voter’s assigned district and the Democratic vote share of her k nearest neighbors, where k is the average number of people assigned to each electoral district. Actual 2014 electoral district boundaries are also included.

The Pennsylvania map indicates a high level of dislocation in the inner suburbs around Pittsburgh (in southwest Pennsylvania). Note that voters in the urban core of

Pittsburgh experience low levels of dislocation. They are overwhelmingly Democratic, and the legislature drew an extremely Democratic urban Pittsburgh district. However, Democrats in Pittsburgh’s inner ring of suburbs experience high rates of dislocation. These are the kinds of neighborhoods in which Justice Roberts seems to indicate that representational rights may have been abridged. There are large, relatively densely populated areas that are extremely Democratic, but the legislature’s redistricting plan in 2012 embedded them in comfortably majority-Republican districts.

It is easy to see that the Pittsburgh metropolitan area could have been carved up in alternative ways that would have dramatically reduced the striking discontinuity in Partisan Dislocation on the edges of districts. It would have been possible to divide the city in a way that included more Democrat-leaning suburbs with Democratic urban neighborhoods. This would have led to two rather than one Pittsburgh-oriented districts, but such an arrangement could still involve relatively compact districts.

In Eastern Pennsylvania, the legislature’s gerrymandering efforts involved the creation of meandering districts that aimed not only to pack Democrats into urban Philadelphia, but also to crack Democratic neighborhoods in the educated suburbs, and to prevent smaller Democratic post-industrial cities from stringing together. Again, we see telltale signs of gerrymandering, such as sharp discontinuities in levels of dislocation at district boundaries, such that members of the party drawing the districts (the Republicans) were far less likely to be dislocated than their opponents.

Figure 4 places this map – with districts devised by Republican lawmakers that were later struck down by the Pennsylvania State Supreme Court – beside the map drawn by a Special Master, Stanford Law Professor Nathaniel Persily, at the Court’s request. As the figure shows, the map drawn by the Special Master shows substantially lower levels of Partisan Dislocation. This illustrates a point we explore more systematically in Section 5: high Partisan Dislocation scores are not just indicative of individually gerrymandered districts. Because they are an indicator of districts that carve up communities in unnatural ways, states with high dislocation scores tend to be ones in which district manipulation has resulted in one party winning a share of seats that is significantly out of line with their overall vote share, even after controlling for the spatial distribution of voters.

District-Level Averages

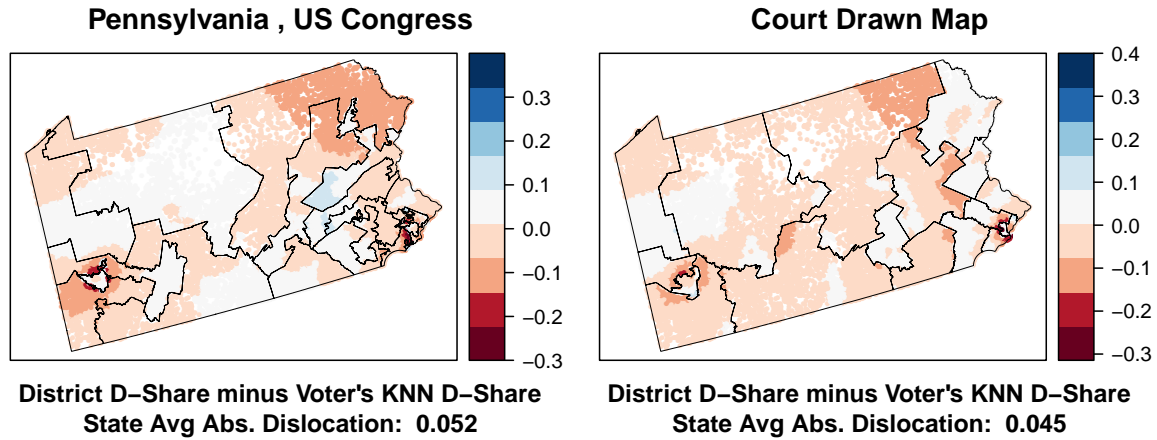
In addition to measuring precinct-level dislocation, we can also aggregate these measures to identify packed and cracked *districts*. In Figure 5, we color districts by their Average Absolute Partisan Dislocation (AAPD) – the average absolute value of representative-voter-level dislocation scores. In particular, the figure again shows the contrast between Pennsylvania’s old maps and those drawn by the Special Master.

The Special Master’s map not only reduces extreme incidences of dislocation around Pittsburgh and in Eastern Pennsylvania, it also reduces *overall* dislocation. By averaging the absolute magnitude of each voter’s dislocation across the entire state, we can get an overall measure of how much an entire map dislocates voters. In the case of

Pennsylvania, for example, we see that the Persily map decreases AAPD by 12.5 % (from 0.052 to 0.045).

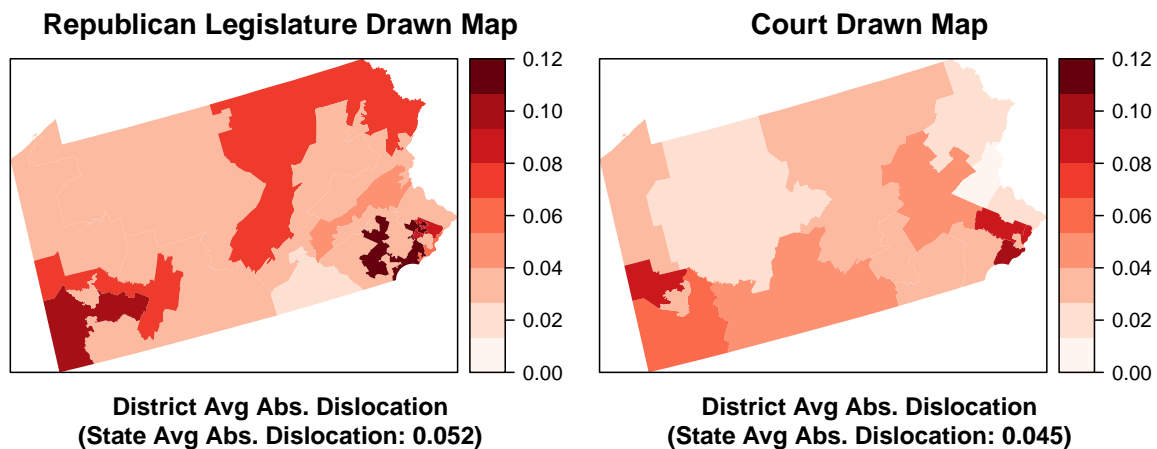
In Figure 6 below, for example, we plot each district's average *absolute* dislocation score. Again, we see that dislocation might be a useful guide to the identification of districts where the notion of local representational harm identified by Justice Roberts is most severe.

Figure 4: Pennsylvania Republican-Drawn and Court-Drawn Districts



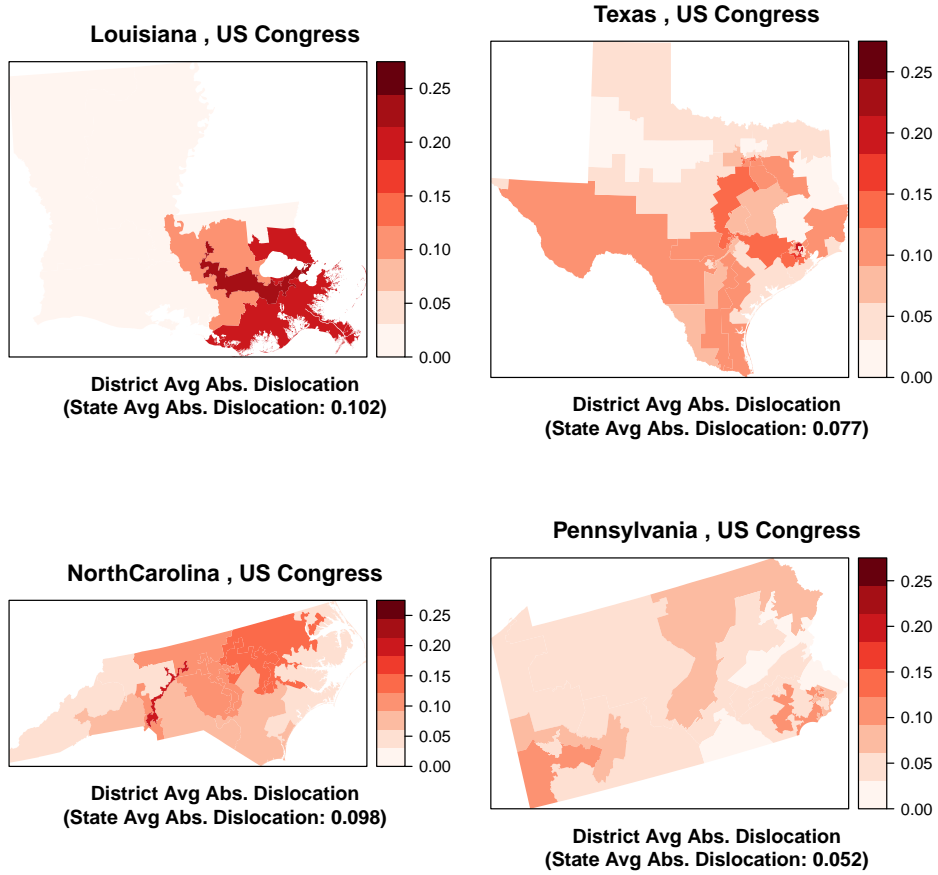
Notes: The above maps plot Partisan Dislocation scores for a set of representative voters. Dislocation is calculated as the difference in the Democratic vote share of each voter's assigned district and the Democratic vote share of her k nearest neighbors, where k is the average number of people assigned to each electoral district. Actual 2014 electoral district boundaries are also included.

Figure 5: Pennsylvania Republican-Drawn and Court-Drawn Districts



Notes: The above maps plot 2014 electoral districts and their AAPD scores. Absolute average dislocation is calculated as the average (over all district voters) of the absolute difference in the Democratic vote share of each voter's assigned district and the Democratic vote share of her k nearest neighbors, where k is the average number of people assigned to each electoral district. District vote shares and the partisanship of nearest neighbors are estimated using precinct-level 2008 US Presidential vote shares as detailed in Section 3.

Figure 6: District AAPD



Notes: The above maps plot 2014 electoral districts and their AAPD scores. Absolute average dislocation is calculated as the average (over all district voters) of the absolute difference in the Democratic vote share of each voter's assigned district and the Democratic vote share of her k nearest neighbors, where k is the average number of people assigned to each electoral district. District vote shares and the partisanship of nearest neighbors are estimated using precinct-level 2008 US Presidential vote shares as detailed in Section 3.

5 Partisan Dislocation and Global Measures of Unfairness and Gerrymandering

As previously noted, one normative basis for concern about gerrymandering is that it generates global representational inequalities. In this section and the one that follows, we will examine the relationship between Political Dislocation and several measures of representational inequality that are often used as metrics of gerrymandering: unequal weighting of votes for one set of partisan voters versus another, a lack of partisan symmetry in seat shares, unusually large seat shares for one party given its electoral geography, and an unusual lack of electoral responsiveness. First, we focus purely on features of enacted plans, and then we focus on measures that require the generation of a large ensemble of alternative plans.

5.1 Dislocation, Votes, and Seats

Some measures of global representational inequality do not rely on comparisons with a sample of non-partisan plans, but rather, calculations based on the distribution of votes and seats across districts in a single enacted plan. One approach, meant to capture whether one party’s voters are relatively more “packed” than those of the other, is to simply calculate the difference between the mean and median of the two-party vote share across districts (McDonald and Best 2015). Another approach, called *Partisan Symmetry* (Katz, King and Rosenblatt 2020), is based on the idea that district maps should generate symmetric conversions from vote shares into seat shares, such that when one party has a 60% vote share, the share of seats they win in the legislature is no different from the number of seats the other party would win with a 60% vote share. Note that Partisan Symmetry implies both parties should win 50% of seats if they have 50% vote shares, but does not imply proportionality, since the seat shares won by parties can take on any value when vote shares deviate from 50% so long as they are symmetric.

To illustrate the ability of Partisan Dislocation to detect these notions of gerrymandering, we first plot the relationship between a state’s AAPD and a set of other metrics for measuring gerrymandering. Before presenting these, however, it is important to emphasize that none of these alternative measures are without their own problems (see Katz, King and Rosenblatt (2020) for extensive discussion of these issues) – indeed, it is precisely because of their limitations that we have developed our Dislocation measure. As such, what we are looking for in these figures is a generally positive relationship, but outliers are to be expected, and as discussed below, often illustrate the value of Partisan Dislocation.

First, in Figure 7, we plot the AAPD score for each enacted districting plan against what is perhaps the simplest global measures of partisan fairness: the mean-median score. The mean-median score is the the absolute value of the difference between the partisanship of the median district and the cross-district mean, calculated using the same vote data employed in our primary analysis (precinct-level returns from the

2008 presidential election). This measure is thought to be instrumentally valuable in detecting gerrymanders that generate unfair seat allocations by packing voters in homogeneous districts, and is also appealing to those interested in partisan symmetry (Best et al. 2018; McDonald and Best 2015). However, because of its myopic focus on only the median district, mean-median scores can fail to identify gerrymandering manipulations in non-median districts, particularly when the statewide partisan baseline is far from 50%. This is related to the issues around responsiveness discussed in Section 5.2, since if the statewide mean is far from 50%, packing and cracking can be employed while maintaining the median district at the mean.⁷ Moreover, unlike our measure, the mean-median score does not take into consideration the political geography of the state or the possible role of the Voting Rights Act.

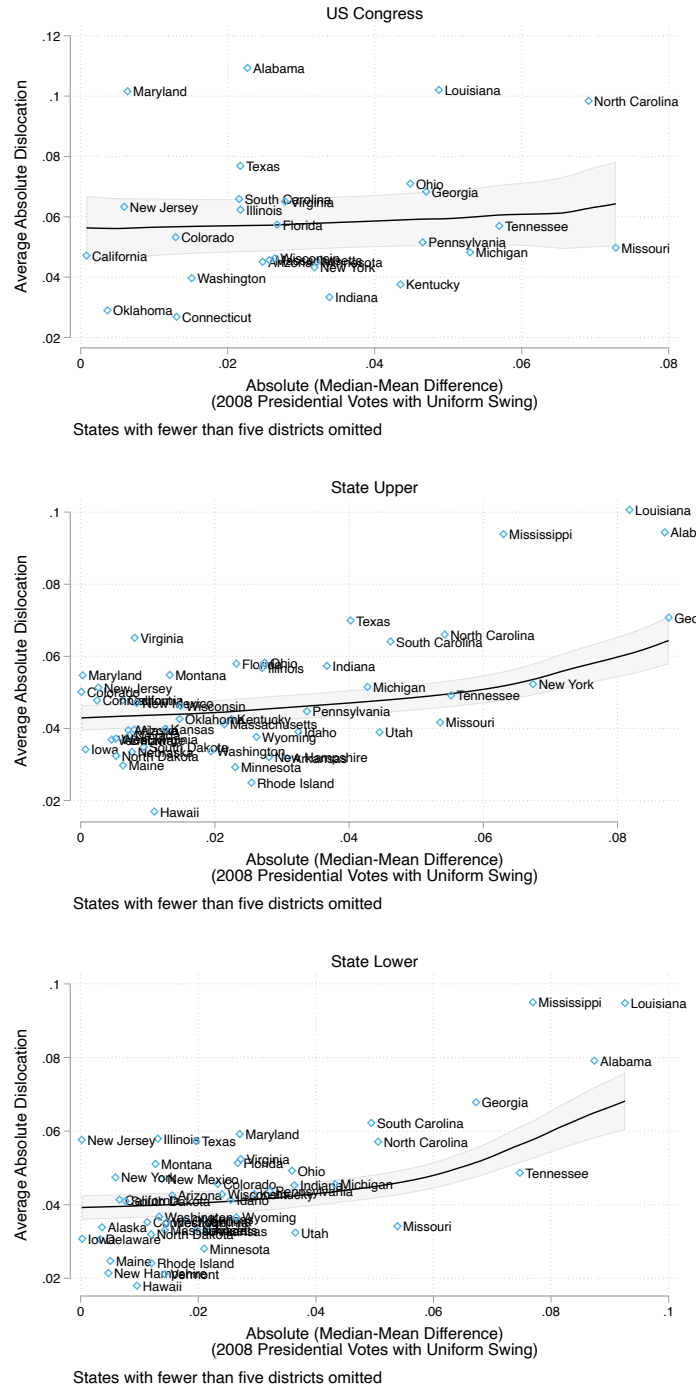
As the Figure shows, while the correlation is not overwhelming, we see that AAPD does tend to track with mean-median scores in the enacted plans. But it is from the exceptions that perhaps we learn the most. For example, consider Texas’ US Congressional districts. As discussed above, the Texas legislature has clearly engaged in gerrymandering, and yet scores low on the mean-median measure. In terms of AAPD, by contrast, Texas scores as the 6th most gerrymandered state in the Union. Similarly Maryland, subject of the recent US Supreme Court gerrymandering case *Benisek v. Lamone*, has a low mean-median score, but the third-highest AAPD score. Missouri demonstrates a high mean-median difference, in part because of the concentration of Democrats in St. Louis and Kansas City and the explicit goal of producing a Congressional district that can be won by minority candidates in both cities. Yet the relatively moderate dislocation value indicates that it was possible to achieve this goal without exceptional levels of dislocation.

In Figure 8, for Congressional districts, we plot AAPD against the Partisan Gini, a measure of the asymmetry of the vote-seat curve (i.e. the degree to which Partisan Symmetry has been violated). This function was introduced as measure 7 in (Grofman 1983) and more recently in (Katz, King and Rosenblatt 2020). An electoral system is said to satisfy the partisan symmetry standard if this value is zero. As the Figure shows, we find a very strong positive relationship between AAPD and the Partisan Gini. That is to say, the *enacted* plans that produce high levels of Partisan Dislocation are also those that produce high levels of partisan asymmetry, such that one party (typically the Republican Party) can expect a relatively high seat share given its vote share.

Finally, as shown in Table 1 below, there is also substantial circumstantial evidence that AAPD scores capture deliberate map manipulation, as AAPD tends to be highest

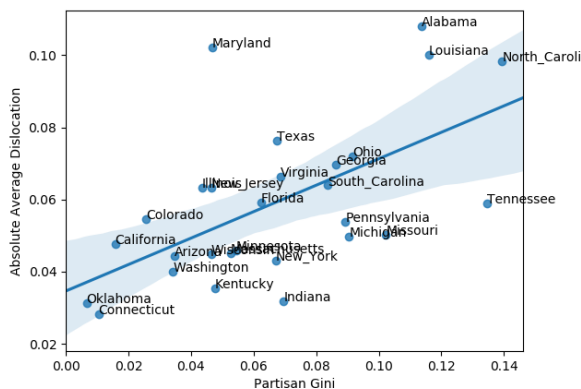
⁷The ensembles used for our analysis below provide evidence for this assertion, as restricting to only plans with absolute mean-median differences less than .01 does not change the range of seat outcomes in the majority of the ensembles, while the majority of the remainder differ only by a single seat outcome. As an example, in Maryland, the ensemble finds maps with between four and seven Democratic seats and maps with each of those seat outcomes can still be found after restricting to the fewer than 20% of plans with mean-median difference less than .01. Similarly, the Texas ensemble with the most constraining VRA bounds finds between six and fourteen seats in the full ensemble, and those seat values also occur in the fewer than 9% of plans with mean-median difference less than .01.

Figure 7: Absolute Average Dislocation and Absolute Median-Mean Scores



Notes: The above figures plot the AAPD score for states (averaged across all voters) against each state's Absolute Mean-Median Difference. Mean-Median Differences are calculated as the absolute difference between the Democratic vote share of the median 2014 district and the average Democratic vote share across all districts. District vote shares and the Partisan Dislocation scores are estimated using precinct-level 2008 US Presidential vote shares as detailed in Section 3. Results are very similar using Democratic vote shares from 2012 to calculate Absolute Mean-Median Differences.

Figure 8: Partisan Gini and AAPD



Notes: The above figure plots AAPD scores for states' enacted 2014 US Congressional district plans against Partisan Gini scores for those same districts. All voting data comes from the 2008 Presidential Election. Shaded regions represent 95% confidence intervals.

in states where district maps were drawn under unified party control.⁸ This is especially true when districts were drawn under unified Republican control, reflecting the success of Republican lawmakers in their efforts to maximize the opportunities presented by redistricting in the early 2010s.

Table 1: State Average Absolute Dislocation by District Creators

	State Lower	State Upper	US House	Overall Avg
Unified Republican Control	0.051	0.055	0.053	0.053
Unified Democratic Control	0.037	0.041	0.045	0.041
Non-Unified or Independent	0.041	0.044	0.037	0.041

5.2 Dislocation and Simulated Districts

The problem with simple statistics generated from the distribution of votes and seats across districts associated with enacted plans, of course, is that they do not take the political geography of the state into account. Nor do they take into account the possibility that unfairness or partisan asymmetry may have been driven, in whole or in part, by efforts to comply with the Voting Rights Act. The solution in the academic literature, and in court, is to compare properties of enacted district maps with ensembles of thousands of sampled maps.

Plaintiffs have argued that voters' rights to equal representation have been violated when gerrymandering results in a party receiving fewer seats than they would absent

⁸Data on who drew districts in each state comes from <http://redistricting.ils.edu/>.

manipulation of district boundaries for political gain. In this framework, parties are not entitled to proportional representation, but through simulation-based methods, plaintiffs attempt to argue that enacted maps result in seat shares that do not arise naturally given the vote shares and the spatial distribution of voters in a state. An additional approach is to contrast the responsiveness of the enacted and simulated maps.

In this section, we augment our analysis by generating 100,000 alternative maps for each state. This allows us to do two useful things. First, we can examine the relationship between our measure of gerrymandering and ensemble-based measures that are increasingly used in court cases. Second, we demonstrate that our measure provides an attractive alternative to anticipated seat shares as a basis for contrasting enacted and sampled plans.

To generate a large collection of comparison plans, we use the ReCom Markov chain introduced in DeFord, Duchin and Solomon (2019) as implemented in the GerryChain software package (MGGG 2019)⁹ to construct ensembles of 100,000 random district plans for each state. The plans generated by the Markov chain are contiguous, population balanced to within 1% of ideal, and are further constrained to preserve the existence of majority-minority districts (to ensure compliance with the Voting Rights Act). Additional details can be found in Appendix D. Note that when calculating seat shares, we add a uniform swing of 3.69% to the two-party vote share of Republicans to bring the overall vote share of Democrats and Republicans to 50-50 nationally. All other measures are uniform-swing invariant.

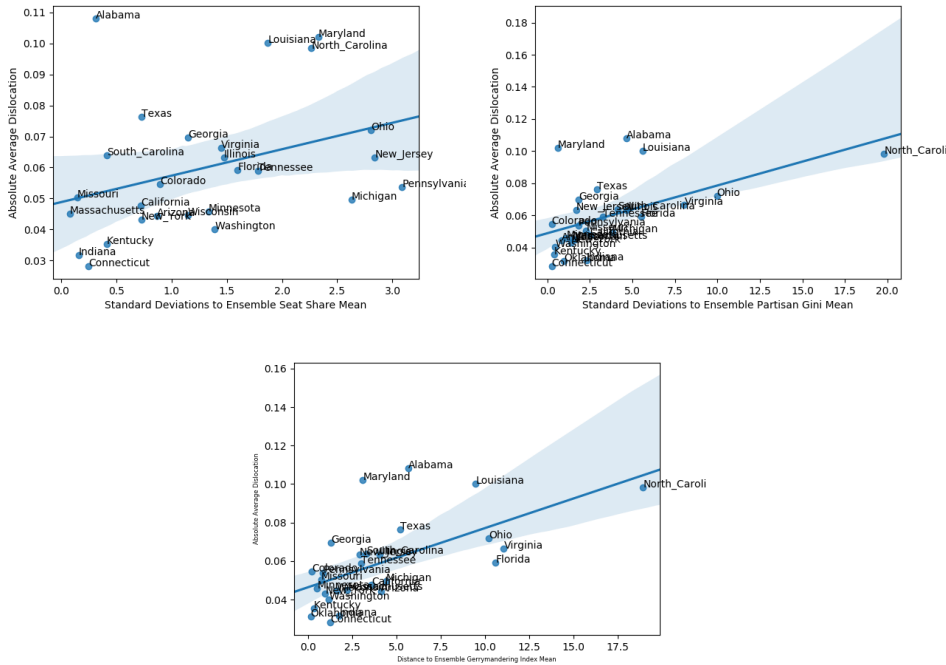
Using these comparison plans, we can now compare AAPD scores to simulation-based metrics of gerrymandering. First, a standard approach is to aggregate precinct-level partisan data to the level of enacted and simulated districts, and examine the difference between the anticipated seat shares for the two parties associated with the enacted plan and those of the simulated plans. Another approach is to examine the difference between the partisan symmetry of the enacted plan and those of the ensemble of simulated plans. A third and more recent approach is to examine the responsiveness of enacted and simulated plans. We measure responsiveness using the *Gerrymandering Index*, based on the work of Herschlag et al. analyzing ensembles of plans in North Carolina and Wisconsin (Herschlag et al. 2018; Herschlag, Ravier and Mattingly 2017). The index is designed to detect maps that create an unusual number of “safe districts” (with, say, a 55% or 60% vote share), and takes on large values when those are present.¹⁰

These clumps of safe districts are often used by gerrymanderers to ensure that seat shares will not respond smoothly to changes in overall vote shares (i.e. seat shares will

⁹<https://github.com/mggg/gerrychain>

¹⁰To calculate the measure, for each plan in the ensemble, we sort the districts by Democratic vote share from smallest to largest, then compute the medians for each ranked position (so the median of the least Democratic districts over all the plans, then the median of the second least Democratic districts, all the way up to the most Democratic favoring). We then calculate the Gerrymandering Index for a given plan by sorting its districts and computing the square root of the sum of the squared differences between the given plan’s values and the corresponding ensemble medians.

Figure 9



Notes: The above figures plot AAPD scores for states’ enacted 2014 US Congressional district plans against simulation-based measures of gerrymandering. Simulation-based measures report the difference between enacted plan scores and the average score across all ensemble plans (in standard deviations of simulated district plans). Figures include only results for states with five or more districts. As detailed in Appendix D, simulated district plans are subject to compactness and population balance constraints, and all plans have the same number of districts that are more than 45% minority (Black or Hispanic) as enacted plans. Shaded regions represent 95% confidence intervals.

not be responsive to changes in vote shares) – instead, because so many districts are stacked with 5-10 percentage point margins, vote swings of less than 5-10% will have no impact on the outcome of elections in those districts, preventing seat shares from responding to changes in vote shares. Note that this metric can only be calculated with the use of simulated ensembles.

Figure 9 below plots AAPD against the distance (in standard deviations) between ensemble average values of various map attributes and those of enacted plans. In particular, the figure compares ensemble and enacted plans in terms of Democratic seat shares, Political Gini, and finally, the Gerrymandering Index described above.

We see a clear positive correlation between our measure of gerrymandering and the simulation-based measure. Note that this relationship is clearly in evidence despite the fact that AAPD scores are *not* normalized against simulation averages – these are raw scores. As shown in Appendix E, correlations are even stronger when we also normalize AAPD scores (by reporting the difference between AAPD for the enacted plan and the average AAPD among simulated districts), but we report the raw scores here to illustrate the value of our measure even absent simulations.

The results also illustrate the limitations of other measures. Seats-based and symmetry-based approaches, for example, tend not to identify Texas as being particularly gerrymandered, while AAPD flags it as a significant gerrymander. Only in the Gerrymandering Index results do we see a simulation-based measure that characterizes Texas as a notable gerrymander: it is the 7th most gerrymandered state in terms of responsiveness dimension, while it is the 5th most according to AAPD.

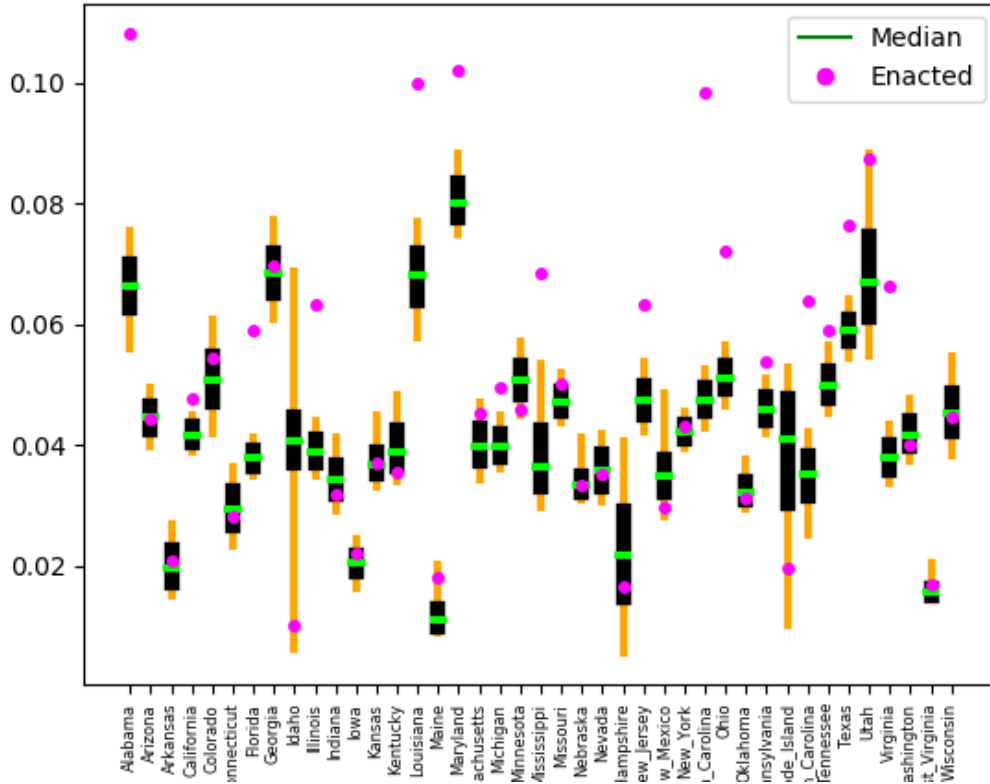
In sum, we view these scatterplots as a validation of our approach. It reaches broadly similar conclusions as existing simulation-based approaches, without requiring a complex computational endeavor that can take considerable time, computing power, and technical expertise. Moreover, areas of disagreement between our approach and existing approaches suggest that our approach can pick up a different class of gerrymanders that might be missed by other approaches.

However, Partisan Dislocation might be most useful not as a substitute, but as a complement to redistricting simulations. As discussed above, it is not always clear which underlying election results should be used in the calculation of hypothetical seat shares, partisan symmetry, or responsiveness. In some situations, the choice can be consequential. Even in an era of nationalized politics, the spatial distribution of voting behavior can vary substantially from one race to another, even for races held on the same day (Rodden and Weighill 2020).

AAPD can also be a useful alternative metric for comparing ensembles of simulated maps with enacted maps. Figure 10 below plots, for each state, the distribution of AAPD score for 100,000 simulated district plans as well as the score for the currently enacted plan. As the Figure clearly shows, it is not just the case that the existing maps of known gerrymanders have high AAPD scores compared to other states, as demonstrated above. They also have much higher AAPD scores than randomly generated districting plans for their own state. Analyses like that contained in Figure 10 might prove to be a very useful diagnostic tool. The outliers in Figure 10 seems to identify all of the well-known gerrymanders of the last redistricting cycle without producing any worrisome false positives.

AAPD might be especially useful in contexts like Louisiana, where states are required to establish majority-minority districts. As noted in Section 4, because Partisan Dislocation scores tend to identify “unnatural” districting plans, there may be occasions where higher-than-usual Dislocation scores are necessary in order to achieve other goals, like facilitating the ability of minorities to elect candidates of choice, if these goals require drawing “unnatural” districts. In states like Alabama, or in Northern Florida, one might need to tolerate a relatively high level of dislocation in order to draw a district where minorities can elect candidates of choice. But Partisan Dislocation also helps measure the *degree* to which communities of like-minded voters have been torn asunder to achieve these ends. Indeed, as we can see in Figure 10, the vast majority of simulated district plans have far lower Partisan Dislocation scores than Louisiana’s enacted plan. The same is true of Alabama, Florida, North Carolina, and Virginia. Recall that in each case, the simulated plans are explicitly drawn so as to provide similar numbers of districts in which minorities can elect candidates of choice as the enacted plans. In

Figure 10: AAPD Scores for Enacted and Random Districts



Notes: The above figure plot AAPD scores for both enacted district plans and the distribution of AAPD scores in simulated district plans. As detailed in Appendix D, simulated district plans are subject to compactness and population balance constraints, and all plans have the same number of districts that are more than 45% minority (Black or Hispanic) as enacted plans.

short, Partisan Dislocation provides a way to identify manipulation above and beyond that which was required in order to comply with the VRA, and to provide detailed maps of precisely where the manipulation took place.

6 Dislocation In Practice: Promise and Peril

Given the properties described here, it is worth pausing to detail some of the applications for which Partisan Dislocation is potentially well suited.

First and foremost, Partisan Dislocation is ideally suited for diagnosing packing and cracking in enacted plans. Not only can AAPD be used to evaluate the overall level of gerrymandering in a state, but (unlike simulations), Partisan Dislocation can also identify specific locations where districts have likely been manipulated for partisan gain. As shown in Figure 10, however, different states have different baseline levels

of Dislocation. As a result, in high-stakes situations like gerrymandering litigation, Partisan Dislocation is probably best analyzed in explicit comparison with simulated ensembles that take political geography and minority representation into account.

In certain situations, Partisan Dislocation can also be used prospectively. Map-drawers interested in developing maps that reflect voter geography closely, in that they tend to be compact and resemble maps near the central tendency of simulated ensembles and which are thus likely to meet current judicial criteria for fairness, and maps that have relatively smooth electoral responsiveness of representation, can make significant progress in this direction by simply working to minimize Partisan Dislocation.

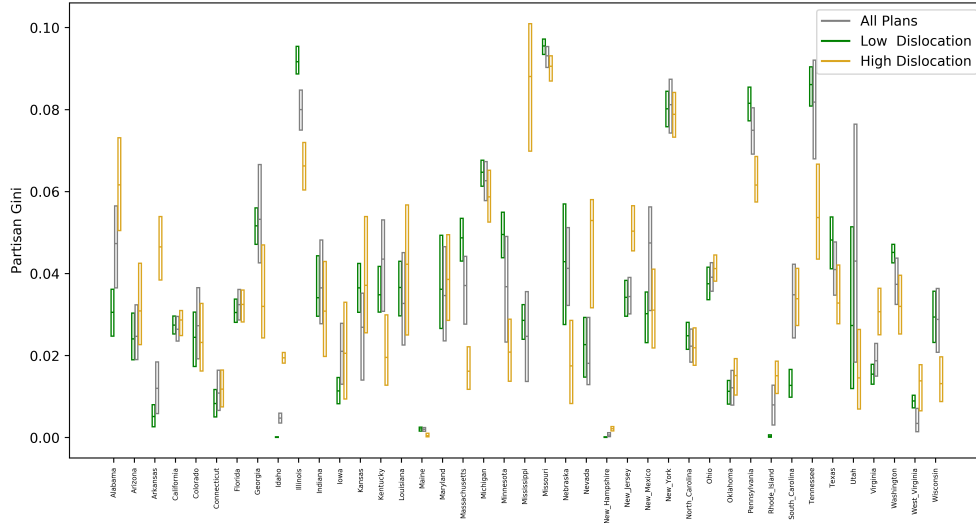
But of course, many do not accept the notion that reflecting voter geography is an appropriate normative goal. Maps with low dislocation scores might accurately reflect voter geography, but they may also serve to entrench the partisan asymmetry that emerges from the clustering of Democratic voters in cities (Rodden 2020). One way to view this residential clustering of Democrats is as a choice made by voters—and thus not a situation requiring accommodation by map makers. But others might argue that low-income voters living in segregated communities reinforced by decades of redlining and racially biased housing policies cannot really be said to have “chosen” to live in communities that tend to result in less efficient representation. In any case, for those who wish to elevate partisan fairness as the primary goal in redistricting, it may sometimes be necessary to break up urban concentrations of Democrats, thus creating plans with high levels of Dislocation.

To illustrate this potential tension between “naturalness” and other normative goals, like partisan symmetry. In Figure 11 we examine the relationship between Partisan Dislocation scores and Partisan Gini. In this Figure, we split our random plan ensembles into the 1% with the highest AAPD, the 1% with the lowest AAPD. We then plot the inter-quartile range of high AAPD plans, low AAPD plans, and the full ensemble of plans in terms of their Partisan Gini scores.

Figure 11 displays interesting heterogeneity across states when it comes to partisan symmetry. In states where Democrats are not highly concentrated in space at the scale of Congressional districts, like Iowa with its dispersed small cities, or Arizona with its vast suburban sprawl, we see that the plans with low and high levels of dislocation are not all that different from the overall distribution in terms of partisan symmetry, and if anything, higher levels of dislocation are associated with higher levels of partisan asymmetry. In these states, the goal of minimizing Dislocation and that of partisan fairness may not be in conflict.

However, the story changes when we look at the early-industrializing states, as well as some Southern states, where Democrats are concentrated in large cities. Examples of the former include Illinois, Michigan, Missouri, Minnesota, Pennsylvania, and Wisconsin. The latter include Georgia, Tennessee, Kentucky, and Texas. In these states, the plans with the highest levels of Dislocation are actually those with the lowest partisan gini. That is to say, they are the plans that minimize the partisan asymmetry associated with the geographic clustering of Democrats. Likewise, the plans with the lowest levels of Dislocation—those that keep local communities of co-partisans together—are

Figure 11



Notes: The above figure plots the inter-quartile range of the 1% of plans with the highest AAPD, the 1% of plans with the lowest AAPD, and the range of all plans in the ensemble in terms of the Partisan Gini score.

the ones that generate the highest levels of partisan asymmetry in the transformation of votes to seats.

This indicates that in a number of populous states, the goals of minimizing Partisan Dislocation and enhancing partisan fairness might be in direct conflict. In order to facilitate partisan symmetry in a state like Illinois—where Democrats are highly concentrated in Chicago— it is necessary to break up communities of co-partisans.

Even when attempting to achieve other goals, like partisan fairness or racial representation, Partisan Dislocation may still be valuable as a tool for ensuring that states enact the *least disruptive* implementation that achieves a given goal. This can be accomplished by requiring states to pick plans that minimize Partisan Dislocation *subject to other fairness or racial representation constraints*. For example, the creation of majority-minority districts has increasingly become a convenient excuse for gerrymandering in many states, as legislators may pack voters strategically in the name of creating these districts. As shown in Louisiana, however, Partisan Dislocation can be used to compare different plans that achieve the same goal, and identify those that seem to achieve those goals in more or less manipulative ways. In doing so, Partisan Dislocation may help limit the space for manipulation in the name of other objectives.

A worthy goal for future research is to generate additional measures, inspired the Dislocation concept, that are less geared toward ex-post gerrymandering detection, and better suited for prospective district-drawers who wish to thread the needle between fairness and “naturalness.” For instance, instead of calculating the continuous difference between the vote share of each representative voter’s neighborhood and that of the

enacted district, one could try to maximize the share of voters with a match between the *binary* partisanship of their neighborhood (is it majority Democratic or Republican?) and that of the enacted district.

Finally, our measure might also be useful in empirical political economy research. Models of distributive politics point to important implications for distributive politics when legislative district lines carve up political communities Stashko (N.d.). For instance, a strategic politician might face incentives to ignore clusters of dislocated members of the minority out-party within a district, or to place unpopular projects, like low-income housing developments or waste processing facilities, in such neighborhoods. If such phenomena are sufficiently pronounced, it is plausible that redistricting would have an impact on property values in dislocated communities.

7 Conclusion

Partisan gerrymandering is difficult to measure, and it is conceptually distinct from partisan fairness, which is typically measured globally rather than locally. It is evident that courts would benefit from a measure that focuses clearly on intentional packing and cracking, rather than fairness, and does so at the level of specific districts or even neighborhoods. We have developed such a measure, called *Partisan Dislocation*, and we have shown that it seems well-suited to the identification of voters that have been cracked or packed. At the level of states, an aggregated measure of dislocation is weakly correlated with global measures of fairness, and more strongly correlated with existing measures of gerrymandering that rely on comparisons of simulated and enacted plans.

Partisan Dislocation might be useful for future litigants wishing to establish that plaintiffs have been directly harmed by being placed in packed or cracked districts. Partisan Dislocation comports with intuitions about how gerrymandering is accomplished, identifies deliberate district manipulations, and if issues of standing (analogous to those raised in *Gill v. Whitford*) arise in state courts, Partisan Dislocation may also be especially helpful. Moreover, it allows for rigorous district-specific gerrymandering analysis.

Partisan Dislocation also holds out promise as a statewide measure of gerrymandering that fills in some of the blind spots of existing approaches, and might serve as a complement to the dominant approach to sampling. In particular, it provides a valuable metric for evaluating a specific plan in relation to a large ensemble of alternative plans. A gerrymandered plan will exhibit significantly higher levels of dislocation than a sample of non-partisan plans, and this analysis can allow for the identification of gerrymandered regions, districts, and even neighborhoods.

Partisan Dislocation also provides a potential template for measuring other forms of dislocation, including racial or economic dislocation. Indeed, the authors are already working to test the usefulness of this approach to identifying efforts to dilute the influence of minority groups in anticipation of redistricting following the 2020 census.

We have drawn a contrast between our approach to gerrymandering, which focuses

purely on the cracking and packing of geographic clusters of partisans, and most existing approaches, which focus on some notion of partisan fairness in the final allocation of seats. There are normative arguments in favor of keeping geographic clusters of like-minded people together in the same district, and of course there are normative arguments in favor of overall fairness in the transformation of votes to seats. Ideally, redistricting plans that achieved one goal would also achieve the other. Unfortunately, we discover that in many but not all U.S. states, the residential geography of Democrats and Republicans might produce a trade-off: districts that minimize Partisan Dislocation will produce asymmetric transformations of votes to seats. This trade-off is among the key challenges facing non-partisan commissioners and court-appointed special masters who hope to achieve both goals.

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A Sampling Variability

As noted in Section 3, our estimates of voter dislocation are subject to two forms of sampling variability: downsampling the number of voters, and then placement of these voters within each precinct.

The first source of variance comes from our need to downsample the universe of all US voters for computational tractability. In particular, we create a set of “representative voters” in each precinct for each party by taking a binomial draw from the total number of actual voters for each party in each precinct. The binomial probability varies by state-chamber, but is equal to $prob_k = \frac{\text{number of districts}}{\text{number of voters in state}} * k$, where $k=1,000$ for state legislative districts and 5,000 for US Congressional districts. This probability generates k voters per district in expectation. This downsampling makes it computationally feasible to calculate the partisan composition each representative voter’s k nearest neighbors. A larger k is used for US Congressional districts as they are much larger with respect to individual precincts, resulting in lower binomial draw probabilities for each precinct, thus increasing sampling variance.

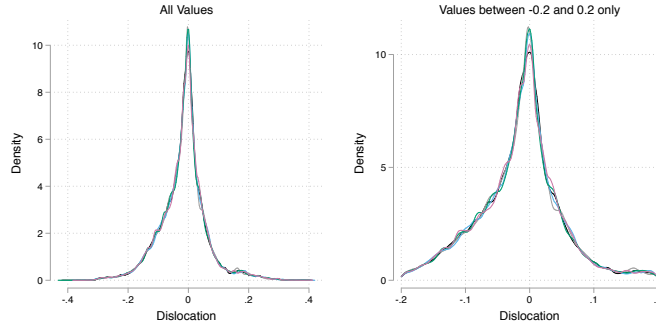
The second source of variance comes from distributing points uniformly within each precinct. Thankfully, US precincts are generally quite geographically compact, limiting the amount of variation introduced by this process.

To evaluate the impact of these sources of variability, Figure 12 below plots the distribution of (representative) precinct-level dislocation scores across five rounds of representative-voter point generation. As the Figures show, variation across each round is extremely small, especially within respect to cross-voter simulation: between-round standard deviations constitute only 0.101 %, 0.103 %, and 0.104 % of total variation for these five rounds for state lower, state upper, and US House chambers respectively.

Figure 13 presents analogous diagnostic distribution at the level of legislative districts (plotting the distribution district-level AAPD scores). Again, between-round standard deviations constitute only 1.11 %, 0.99 %, and 1.67 % of total variation for these five rounds for state lower, state upper, and US representative chambers respectively.

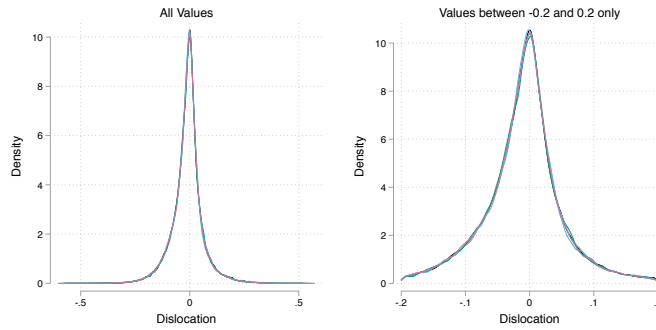
Figure 12

Voter-Level Dislocation Distributions, US Congress
Across 5 Generations of Representative Points



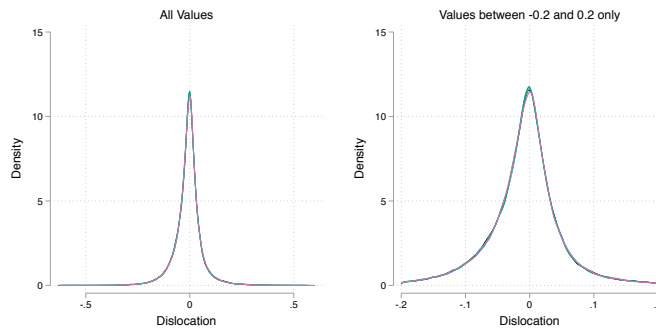
Within Simulation Std. Dev.: 0.0776, Between Simulation Std. Dev.: 0.0001
Between As Pct of Total Std. Dev.: 0.104%
Kernel densities plotted from 10% sample; variance decomposition from full sample.

Voter-Level Dislocation Distributions, State Upper
Across 5 Generations of Representative Points



Within Simulation Std. Dev.: 0.0709, Between Simulation Std. Dev.: 0.0001
Between As Pct of Total Std. Dev.: 0.103%
Kernel densities plotted from 10% sample; variance decomposition from full sample.

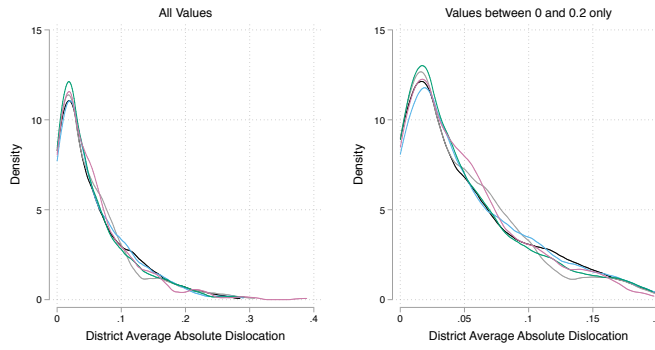
Voter-Level Dislocation Distributions, State Lower
Across 5 Generations of Representative Points



Within Simulation Std. Dev.: 0.0659, Between Simulation Std. Dev.: 0.0001
Between As Pct of Total Std. Dev.: 0.101%
Kernel densities plotted from 10% sample; variance decomposition from full sample.

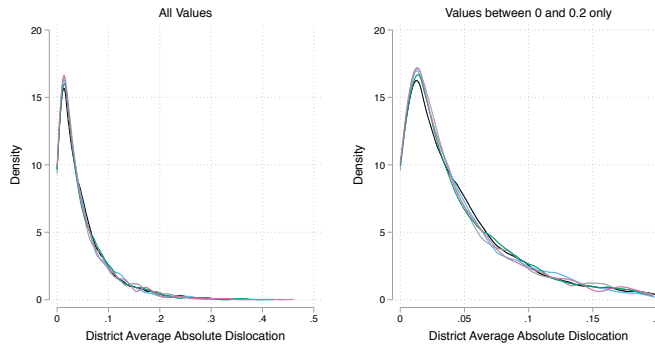
Figure 13

District-Level Absolute Dislocation Distributions, US Congress
Across 5 Generations of Representative Points



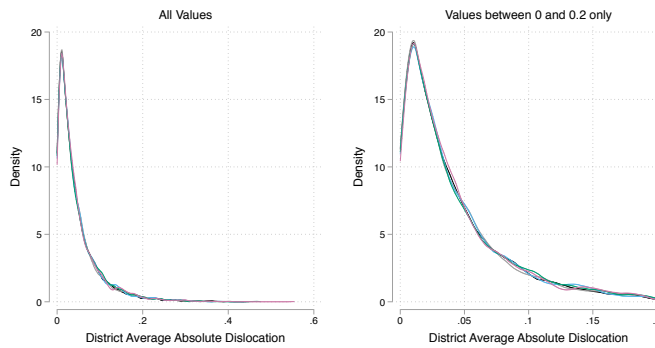
Within Simulation Std. Dev.: 0.0558, Between Simulation Std. Dev.: 0.0009
Between As Pct of Total Std. Dev.: 1.67%

District-Level Absolute Dislocation Distributions, State Upper
Across 5 Generations of Representative Points



Within Simulation Std. Dev.: 0.0520, Between Simulation Std. Dev.: 0.0005
Between As Pct of Total Std. Dev.: 0.99%

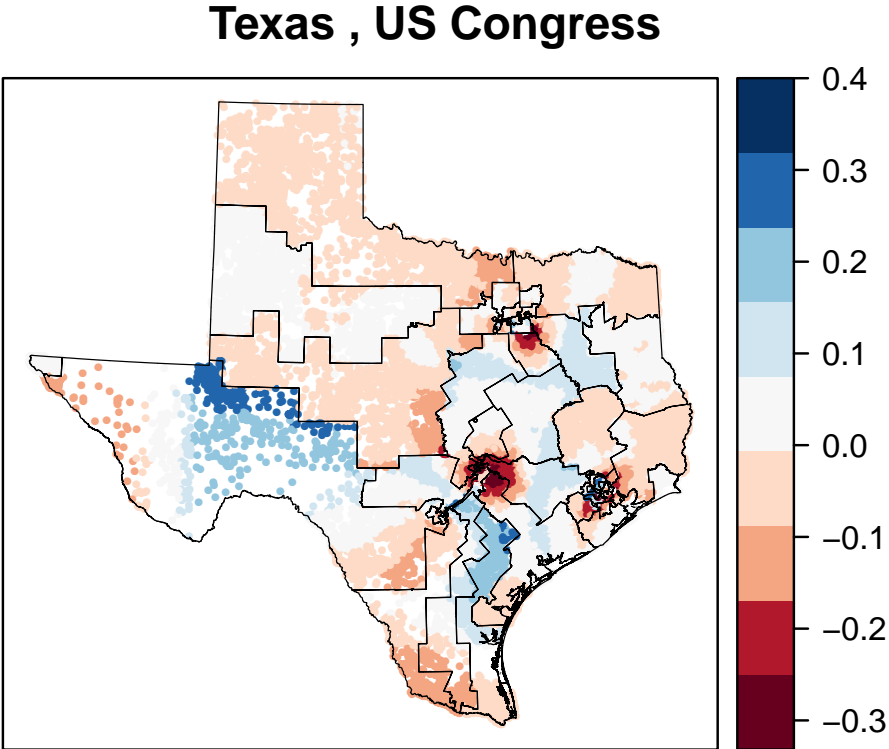
District-Level Absolute Dislocation Distributions, State Lower
Across 5 Generations of Representative Points



Within Simulation Std. Dev.: 0.0512, Between Simulation Std. Dev.: 0.0006
Between As Pct of Total Std. Dev.: 1.11%

B Additional Partisan Dislocation Maps

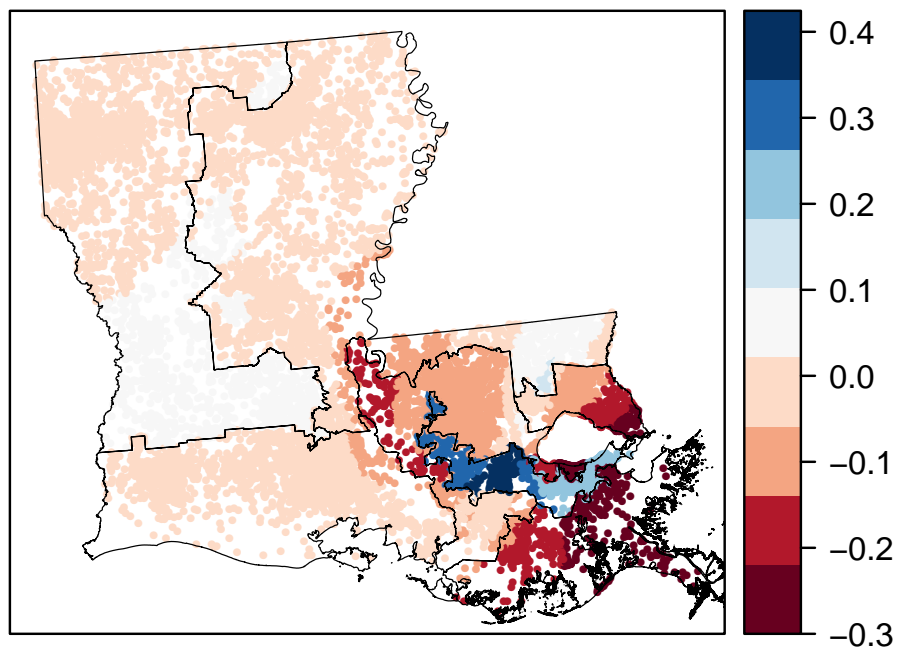
Figure 14: Partisan Dislocation in Texas US House Districts



District D–Share minus Voter's KNN D–Share
State Avg Abs. Dislocation: 0.077

Figure 15: Partisan Dislocation in Louisiana US House Districts

Louisiana , US Congress



District D–Share minus Voter's KNN D–Share
State Avg Abs. Dislocation: 0.102

Figure 16: Partisan Dislocation in North Carolina US House Districts

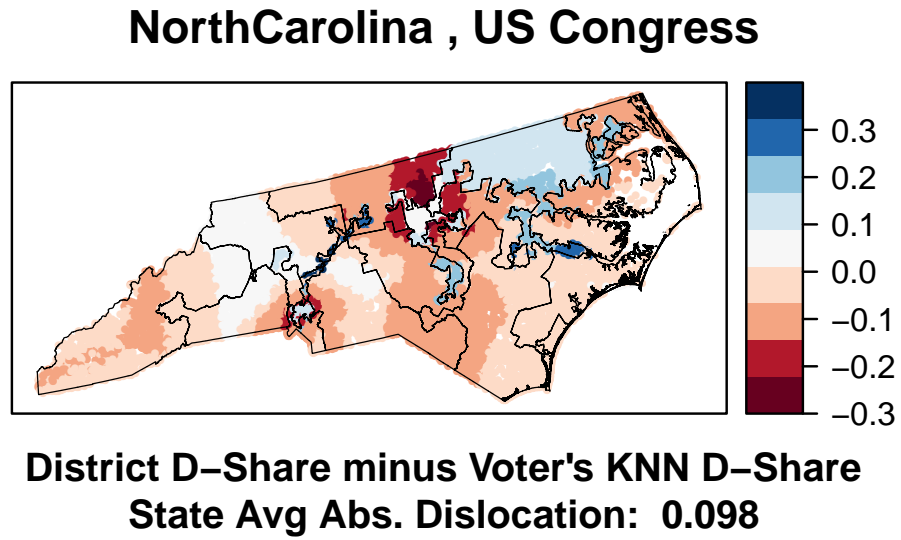
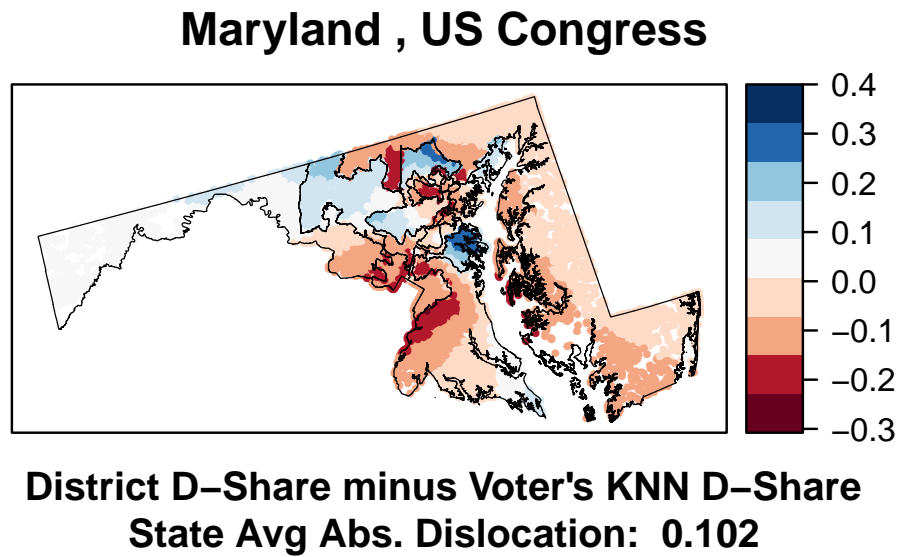


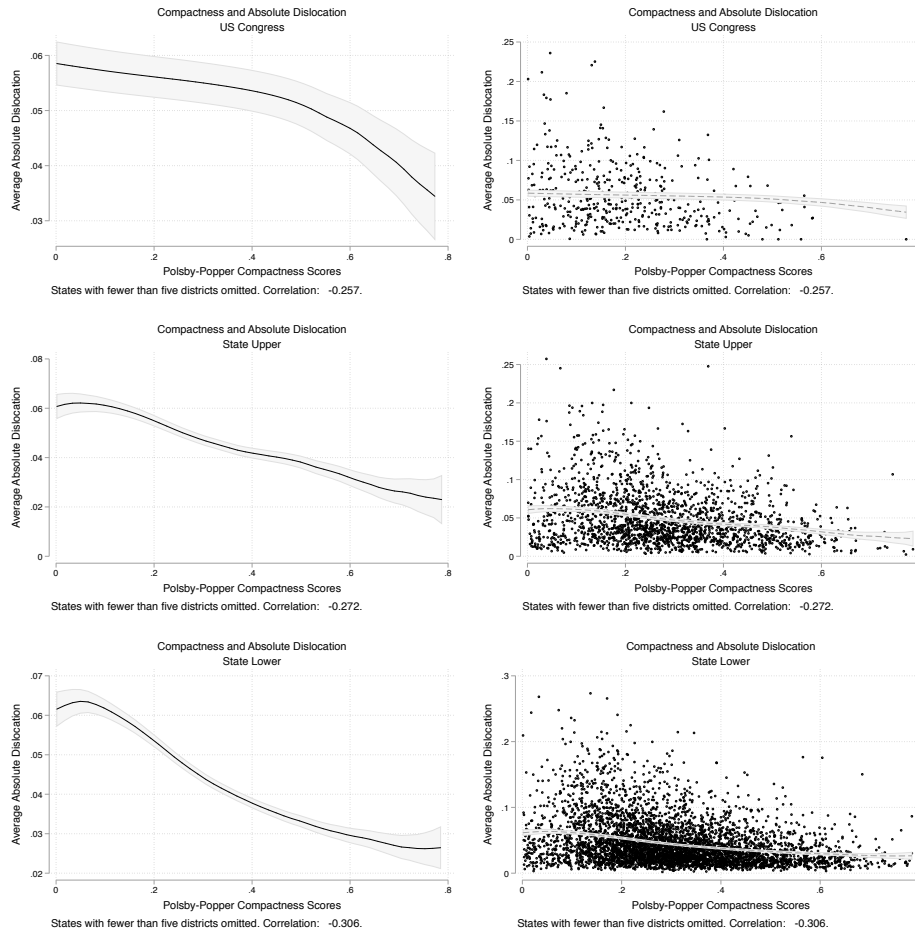
Figure 17: Partisan Dislocation in Maryland US House Districts



C Partisan Dislocation and Compactness

As Partisan Dislocation contrasts the partisan composition of a voter's actual district to what would be the composition of a perfectly compact (circular, modulo boundary reflections) district centered on the voter, one might worry that dislocation simply measures deviations from compactness. As shown in Figure 18 below, while it is the case that dislocation and compactness are related (as we would expect, given the types of deliberately gerrymandered districts dislocation aims to identify) the relationship between the two factors is weak: the correlation is only around ~ -0.275 at all district levels.

Figure 18: District AAPD and District Compactness
With and Without Scatter Overlay



D Simulated Districts

Markov chain based approaches have become a common tool for generating large collections of districting plans. For this analysis, we used the ReCom chain introduced in (DeFord, Duchin and Solomon 2019), which modifies the plan at each step by selecting a pair of adjacent districts, forming a uniform spanning tree on the nodes assigned to those districts, and then selecting a uniformly chosen edge to cut that leaves the remaining parts population balanced to within 1% of ideal. We select the districts to merge proportional to the number of edges on their boundary, in order to promote compactness.

Dual graphs for each state were constructed directly from the precinct shapefiles used in the main analysis. We connected islands and other disconnected regions automatically, finding the nearest precincts in the main body of the state. Florida and California required extra processing, as the shapefiles contained empty polygons that spanned large regions of the state. These outliers were removed and the resulting dual graphs were reconnected.

Initial seeds for the ensembles were constructed using a recursive spanning tree method that generates a single district at a time by drawing a spanning tree for the remaining portion of the dual graph and separating a single edge whose smaller part has population within 1% of ideal. Once the initial population balanced seeds were constructed, an optimization version of the ReCom chain was used to generate starting plans that complied with our chosen VRA bounds.

D.1 Voting Rights Act

In order to model potential impacts of including Voting Rights Act districts in the ensembles, we count the number of districts in each state’s 2014 plans whose adult voting age population is at least 40%, 45%, or 50% Black or Hispanic using data from the 2010 census. We then ensure that all simulated district plans have at least the same number of districts that clear these bars. Although the Voting Rights Act does not necessarily support specific numerical percentages by matching the values observed in the enacted plans we are attempting to generate ensembles that represent similar constraints. Results presented in the paper use a 45% threshold, but our results are similar using 40% or 50%.¹¹ Note that as currently jurisprudence considers the proportion of voting age population in a district that are Black, or the proportion of the voting age population in a district that are Hispanic, but not the proportion of the voting age population that is either Black or Hispanic (so called “coalition” districts), we also use this operationalization.

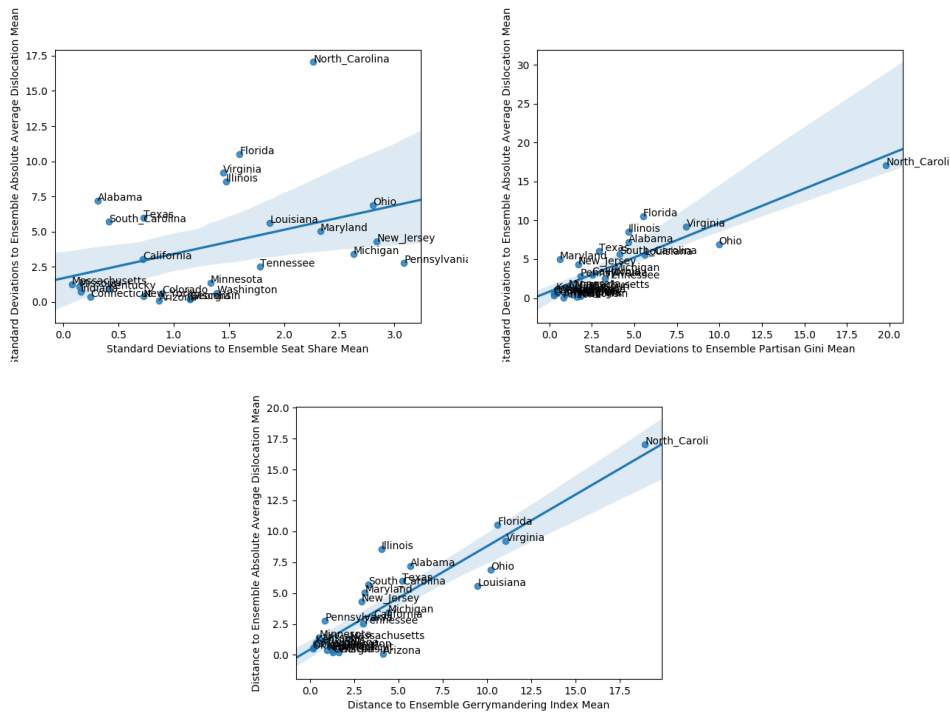
The only exceptions to this procedure were North Carolina and Florida. The map that was in place in North Carolina in 2014 was ruled unconstitutional as a racially

¹¹Despite the term “majority-minority,” it is rarely the case that the majority of voting age populations in majority-minority districts are actually minority. Exact thresholds vary across states and court cases, however.

packed gerrymander and matching the percentages from that plan would have encoded this packing. Instead, the three ensembles kept two districts over 40%, one district over 40% and one district over 35%, and one district over 40%, respectively. The plots in the main text use the middle ensemble, which is very similar to the approach used in (Herschlag et al. 2018) and related expert testimony in court. In Florida, the state of the precinct data discussed above made it difficult to match the values observed in the enacted plan for Black percentage districts. Thus, districts we used we used bounds of two districts over 40% and one over 35%, one over 45% and one over 35%, and one over 50%, respectively, while computing the Hispanic district bounds as in the other states.

E Simulation-Based Metrics and Simulation-Normalized AAPD

Figure 19



Notes: The above figures plot normalized AAPD scores for states' enacted 2014 US Congressional district plans against simulation-based measures of gerrymandering. Both AAPD and other metrics are normalized by calculating the difference between enacted plan scores and the average score across all ensemble plans (in standard deviations of simulated district plans). Figures include only results for states with five or more districts. As detailed in Appendix D, simulated district plans are subject to compactness and population balance constraints, and all plans have the same number of districts that are more than 45% minority (Black or Hispanic) as enacted plans. Results are similar using either 40% or 50% thresholds for minority share.